Contraction Type Mapping on 2-Metric Spaces

Dhruba Das¹, Kanak Das²

¹(Department of Statistics, Gauhati University, India) ²(Department of Botany, Bajali College, India)

Abstract: The superimposition of infinite number of intervals $[a_1, b_1]$, $[a_2, b_2]$, $[a_3, b_3]$,...., $[a_n, b_n]$ follows two laws of randomness if

(a) $a_i \neq a_j$; i,j = 1, 2,..., n, (b) $b_i \neq b_j$; i,j = 1, 2,..., n, (c) $max(a_i) \leq min(b_i)$; i = 1, 2,..., n, where $n \rightarrow \infty$ **Keywords:** Superimposition of sets, Probability distribution function, Glivenko – Cantelli Lemma.

I. Introduction

Construction of normal fuzzy number has been discussed in ([1], [2]) based on the randomness – fuzziness consistency principle deduced by Baruah ([3], [4], [5]). Based on this aforesaid principle by including two more conditions which are not mentioned by Baruah, we have shown that if we superimpose infinite number of intervals [a₁, b₁], [a₂, b₂], [a₃, b₃],....., [a_n, b_n], then the values $a_{(1)}$, $a_{(2)}$,, $a_{(n)}$ follows an uniform probability distribution function and the values $b_{(1)}$, $b_{(2)}$,...., $b_{(n)}$ follows an another complementary uniform probability distribution function where $a_{(1)}$, $a_{(2)}$,, $a_{(n)}$ and $b_{(1)}$, $b_{(2)}$,, $b_{(n)}$ are arranged in increasing order of magnitude of a_1 , a_2 , a_3 ,, a_n and b_1 , b_2 , b_3 ,, b_n respectively. If $\alpha = \min(a_i)$, $\beta = \max(a_i)$, $\mu = \min(b_i)$, $\gamma = \max(b_i)$, by satisfying the condition $a_i \neq a_j$, $b_i \neq b_j$ and $\max(a_i) < \min(b_i)$; i, j = 1, 2,...., n, then we can define the function $\psi(x)$ as

$$\begin{split} \psi(x) = \psi_1(x) & \text{if } \alpha \le x \le \beta, \\ = 1 - \psi_2(x) & \text{if } \mu \le x \le \gamma, \\ = 1 & \text{if } \beta \le x \le \mu, \\ = 0 & \text{otherwise.} \end{split}$$

Where $\Psi_1(x)$ being a continuous distribution function in the interval $[\alpha, \beta]$, and $(1 - \Psi_2(x))$ being a continuous distribution function in the interval $[\mu, \gamma]$, with $\Psi_1(\alpha) = \Psi_2(\gamma) = 0$ and $\Psi_1(\beta) = \Psi_2(\mu) = 1$.

II. The Operation Of Set Superimposition

The operation of set superimposition of two real intervals $[a_1, b_1]$ and $[a_2, b_2]$ as

$$[a_1, b_1](S)[a_2, b_2] = [a_{(1)}, a_{(2)}] \cup [a_{(2)}, b_{(1)}]^{(2)} \cup [b_{(1)}, b_{(2)}]$$

Where $a_{(1)} = \min(a_1, a_2), a_{(2)} = \max(a_1, a_2), b_{(1)} = \min(b_1, b_2)$ and $b_{(2)} = \max(b_1, b_2)$. Here we have assumed without any loss of generality that $a_1 \neq a_2$, $b_1 \neq b_2$ and $[a_1, b_1] \cap [a_2, b_2]$ is not void or in other words that $\max(a_i) < \min(b_i), i = 1, 2$



Figure 1: Superimposition of $[x_1, y_1]^{\left(\frac{1}{3}\right)}, [x_2, y_2]^{\left(\frac{1}{3}\right)}$ and $[x_3, y_3]^{\left(\frac{1}{3}\right)}$



Figure2: Cumulative and complementary cumulative distribution functions

For the three intervals $[x_1, y_1]^{\binom{1}{3}}, [x_2, y_2]^{\binom{1}{3}}$ and $[x_3, y_3]^{\binom{1}{3}}$ all with elements with a constant probability equal to 1/3 everywhere, we shall have

$$[x_{1}, y_{1}]^{\binom{1}{3}}(S)[x_{2}y_{2}]^{\binom{1}{3}}(S)[x_{3}, y_{3}]^{\binom{1}{3}} = [x_{(1)}, x_{(2)}]^{\binom{1}{3}} \cup [x_{(2)}, x_{(3)}]^{\binom{1}{3}} \cup [x_{(3)}, y_{(1)}]^{(1)} \cup [y_{(1)}, y_{(2)}]^{\binom{1}{3}} \cup [y_{(2)}, y_{(3)}]^{\binom{1}{3}} \cup [y_{(2)}, y_{(2)}]^{\binom{1}{3}} \cup [y_{(2)}, y_{(2$$

where, for example $[y_{(1)}, y_{(2)}]^{2/3}$ represents the interval $[y_{(1)}, y_{(2)}]$ with probability 2/3 for all elements in the entire interval, $x_{(1)}, x_{(2)}, x_{(3)}$ being values of x_1, x_2, x_3 arranged in increasing order of magnitude, and similarly $y_{(1)}, y_{(2)}, y_{(3)}$ being values of y_1, y_2, y_3 arranged in increasing order of magnitude again. We here presumed that $[x_1, y_1] \cap [x_2, y_2] \cap [x_3, y_3]$ is not void and $x_1 \neq x_2 \neq x_3$ and $y_1 \neq y_2 \neq y_3$. It can be seen that for n intervals $[a_1, b_1]^{1/n}$, $[a_2, b_2]^{1/n}$,, $[a_n, b_n]^{1/n}$ all with probability equal to 1/n

everywhere, we shall have

$$\begin{bmatrix} a_1, b_1 \end{bmatrix}_n^{\frac{1}{n}} (S) \begin{bmatrix} a_2, b_2 \end{bmatrix}_n^{\frac{1}{n}} (S) \dots \begin{bmatrix} a_n, b_n \end{bmatrix}_n^{\frac{1}{n}} = \begin{bmatrix} a_{(1)}, a_{(2)} \end{bmatrix}_n^{\frac{1}{n}} \cup \begin{bmatrix} a_{(2)}, a_{(3)} \end{bmatrix}_n^{\frac{2}{n}} \cup \dots \bigcup \begin{bmatrix} a_{(n-1)}, a_{(n)} \end{bmatrix}_n^{\frac{n-1}{n}} \cup \begin{bmatrix} a_{(n-1)}, b_{(n-1)} \end{bmatrix}_n^{\frac{2}{n}} \cup \begin{bmatrix} b_{(n-1)}, b_{(n)} \end{bmatrix}_n^{\frac{1}{n}},$$

Where, for example, $[b_{(1)}, b_{(2)}]^{\frac{n-1}{n}}$ represents the interval $[b_{(1)}, b_{(2)}]$ with probability $\frac{n-1}{n}$ in the entire interval, a₍₁₎, a₍₂₎,....., a_(n) being values of a₁, a_{2,.....}a_n arranged in increasing order of magnitude, and b₍₁₎, b(2),....., b(n) being values of b1, b2,.....bn arranged in increasing order of magnitude. Thus for the intervals $[a_1,b_1]^{\frac{1}{n}}, [a_2,b_2]^{\frac{1}{n}}, \dots, [a_n,b_n]^{\frac{1}{n}}, \text{ all with uniform probability } \frac{1}{n}, \text{ the probabilities of the superimposed}$

intervals are $\frac{1}{n}, \frac{2}{n}, \dots, \frac{n-1}{n}, 1, \frac{n-1}{n}, \dots, \frac{2}{n}$ and $\frac{1}{n}$. These probabilities considered in two halves as

$$\left(0,\frac{1}{n},\frac{2}{n},\ldots,\frac{n-1}{n},1\right)$$

and

$$\left(1,\frac{n-1}{n},\ldots,\frac{2}{n},\frac{1}{n},0\right)$$

would suggest that they can define an empirical distribution and a complementary empirical distribution on a_1, a_2, \dots, a_n and b_1, b_2, \dots, b_n respectively. In other words, for realizations of the values of $a_{(1)}$, $a_{(2)},\ldots,a_{(n)}$ in increasing order and of $b_{(1)}, b_{(2)},\ldots,b_{(n)}$ again in increasing order, we can see that if we define

$$\psi_{1}(x) = \begin{cases} 0 & \text{if } x < a_{(1)} \\ \frac{r-1}{n} & \text{if } a_{(r-1)} \le x \le a_{(r)}, r = 2, 3, ..., n \\ 1 & \text{if } x \ge a_{(n)}, \end{cases}$$

$$\psi_{2}(x) = \begin{cases} 1 & \text{if } x < b_{(1)} \\ 1 - \frac{r-1}{n} & \text{if } b_{(r-1)} \le x \le b_{(r)}, r = 2, 3, ..., n \\ 0 & \text{if } x \ge b_{(n)}, \end{cases}$$

Then the Glivenko - Cantelli Lemma on Order Statistics assures that

 $\psi_1(x) \to \prod_1 [\alpha, x], \qquad \alpha \le x \le \beta, \\ \psi_2(x) \to 1 - \prod_2 [\mu, x], \quad \mu \le x \le \gamma,$

where $\prod_{i} [\alpha, x]$, $\alpha \le x \le \beta$ and $\psi_2(x)$, $\mu \le x \le \gamma$ are two probability distributions. Here in this case we have considered that $\max(a_i) \square \min(b_i)$, but for large number of observation when $\max(a_i) = \min(b_i)$ that is $\beta = \mu$ then $a_{(n)} = b_{(1)}$ and we can write

$$\psi_1(x) \to \prod_1 [\alpha, x], \qquad \alpha \le x \le \beta, \\ \psi_2(x) \to 1 - \prod_2 [\beta, x], \quad \beta \le x \le \gamma,$$

where $\prod_{1} [\alpha, x]$, $\alpha \le x \le \beta$ and $\psi_2(x)$, $\beta \le x \le \gamma$ are two probability distributions.

III. CONCLUSION

The superimposition of an infinite number of intervals $[a_1, b_1]$, $[a_2, b_2]$, $[a_3, b_3]$,....., $[a_n, b_n]$ by satisfying the conditions $a_i \neq a_j$, $b_i \neq b_j$ and $\max(a_i) \leq \min(b_i)$; i,j=1, 2,..., n, follows two laws of randomness, one of which is $a_{(1)}, a_{(2)}, ..., a_{(n)}$ follows an uniform probability distribution function and the other one is $b_{(1)}$, $b_{(2)}$,..., $b_{(n)}$ follows an another complementary uniform probability distribution function where $a_{(1)}, a_{(2)}, ..., a_{(n)}$ and $b_{(1)}, b_{(2)}, ..., b_{(n)}$ are arranged in increasing order of magnitude of $a_1, a_2, a_3, ..., a_n$ and $b_1, b_2, b_3, ..., b_n$ respectively. If $\alpha = \min(a_i), \beta = \max(a_i), \mu = \min(b_i), \gamma = \max(b_i)$ and $\max(a_i) \square \min(b)$, then we can define the function $\psi(x)$ as

$$\psi(x) = \psi_1(x) \quad \text{if } \alpha \le x \le \beta,$$

=1-\psi_2(x) \text{ if } \mu \le x \le \gamma,
=1 \text{ if } \beta \le x \le \mu,
=0 \text{ otherwise.}

Where $\Psi_1(x)$ being a continuous distribution function in the interval $[\alpha, \beta]$, and $(1 - \Psi_2(x))$ being a continuous distribution function in the interval $[\mu, \gamma]$, with $\Psi_1(\alpha) = \Psi_2(\gamma) = 0$ and $\Psi_1(\beta) = \Psi_2(\mu) = 1$.

Again if max(a_i) = min(b_i), then $\beta = \mu$ and we can define the function $\psi(x)$ as

$$\psi(x) = \psi_1(x) \quad \text{if } \alpha \le x \le \beta,$$

=1-\psi_2(x) \text{ if } \beta \le x \le \gamma,
=0 \quad \text{otherwise.}

Where $\Psi_1(x)$ being a continuous distribution function in the interval $[\alpha, \beta]$, and $(1 - \Psi_2(x))$ being a continuous distribution function in the interval $[\beta, \gamma]$, with $\Psi_1(\alpha) = \Psi_2(\gamma) = 0$ and $\Psi_1(\beta) = \Psi_2(\beta) = 1$.

References

- Hemanta K. Baruah, Construction of Normal Fuzzy Numbers Using the Mathematics of Partial Presence, Journal of Modern Mathematics Frontier, Vol. 1, No. 1, 2012, 9 – 15
- [2] Hemanta K. Baruah, The Randomness–Fuzziness Consistency Principle, International Journal of Energy, Information and Communications, Vol. 1, Issue 1, 2010, 37 48
- [3] Hemanta K. Baruah, The Theory of Fuzzy Sets: Beliefs and Realities, *International Journal of Energy Information and Communications*, Vol. 2, Issue 2, 2011, 1 22
- [4] Hemanta K. Baruah, In Search of the Root of Fuzziness: The Measure Theoretic Meaning of Partial Presence, Annals of Fuzzy Mathematics and Informatics, Vol. 2, No. 1, 2011, 57 – 68
- [5] Hemanta K. Baruah, Set Superimposition and Its Applications to the Theory of Fuzzy Sets, Journal of the Assam Science Society, Vol. 40, No. 1 & 2, 1999, 25-31